

Online Appendix for “Ambiguity, Information Processing, and Financial Intermediation”

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Online Appendix A: Comparative Statics Analysis

The risk premium. Given that

$$\pi_{R,t}^U = \frac{\sigma^2(1 + \theta_1/\rho)(1 + \theta_1^h/\rho^h)}{\rho^2 x_t [1 + (\rho^h - \rho)x_t]} \frac{\left[(\rho^h - \rho) + (\theta_1^h - \theta_1) + (1 + \theta_1/\rho) \frac{1}{x_t} \right]}{\left[(1 + \theta_1^h/\rho^h) - (1 + \theta_1/\rho) + (1 + \theta_1/\rho) \frac{1}{\rho x_t} \right]^2},$$

we have

$$\frac{d\pi_{R,t}^U}{d\theta_1} = \frac{2\sigma^2 (\gamma^h)^2 [\rho^h + \theta_1^h + \gamma(1/x_t - \rho)]}{\rho^3 x_t [1 + (\rho^h - \rho)x_t] [\gamma^h - \gamma + \gamma/(\rho x_t)]^3}.$$

To evaluate the sign of this expression, we consider the following possibilities. Possibility (1): If $\gamma^h > \gamma \geq 1$, we have

$$x_t > 0 > -\frac{\gamma}{\rho(\gamma^h - \gamma)} \Leftrightarrow \gamma^h - \gamma + \gamma \frac{1}{\rho x_t} > 0.$$

Possibility (2): If $\gamma > \gamma^h \geq 1$,

$$\frac{\gamma}{\gamma - \gamma^h} > 1 \Leftrightarrow \frac{\gamma}{\rho(\gamma - \gamma^h)} > \frac{1}{\rho} \geq x_t \Leftrightarrow \gamma^h - \gamma + \gamma \frac{1}{\rho x_t} > 0.$$

Possibility (3): If $\gamma^h = \gamma \geq 1$,

$$\gamma^h - \gamma + \gamma \frac{1}{\rho x_t} = \gamma \frac{1}{\rho x_t} > 0.$$

We thus have:

$$\frac{d\pi_{R,t}^U}{d\theta_1} = \frac{2\sigma^2 (\gamma^h)^2 [\rho^h + \theta_1^h + \gamma(1/x_t - \rho)]}{\rho^3 x_t [1 + \Delta\rho x_t] [\gamma^h - \gamma + \gamma/(\rho x_t)]^3} > 0.$$

Given that

$$\pi_{R,t}^U = \frac{1}{P/D} \frac{\sigma^2 (1 + \theta_1/\rho) \beta_t^U}{x_t [\rho^h - (\rho^h - \rho) \beta_t^U]^2},$$

we have

$$\begin{aligned} \frac{d\pi_{R,t}^U}{d\theta_1^h} &= \frac{1}{P/D} \frac{\sigma^2 (1 + \theta_1/\rho)}{x_t [\rho^h - (\rho^h - \rho) \beta_t^U]^3} \left[\frac{d\beta_t^U}{d\theta_1^h} (\rho^h - (\rho^h - \rho) \beta_t^U) + 2 (\rho^h - \rho) \frac{d\beta_t^U}{d\theta_1^h} \beta_t^U \right] \\ &= \frac{\rho^h}{[1 + (\rho^h - \rho) x_t] x_t} \frac{\sigma^2 (1 + \theta_1/\rho)^2 (1/x_t - \rho) (\rho^h + (\rho^h - \rho) \beta_t^U) (\beta_t^U)^2}{(\rho^h - (\rho^h - \rho) \beta_t^U)^3 \rho^h (1 + \theta_1^h/\rho^h)}. \end{aligned}$$

Since $0 \leq \beta_t^U \leq 1$, $\rho \leq \rho^h - (\rho^h - \rho) \beta_t^U \leq \rho^h$ and

$$\frac{d\pi_{R,t}^U}{d\theta_1^h} = \frac{\sigma^2 \gamma^2 (1/x_t - \rho) (\rho^h + \Delta\rho \beta_t^U) (\beta_t^U)^2}{\gamma^h x_t (1 + \Delta\rho x_t) (\rho^h - \Delta\rho \beta_t^U)^3} \geq 0.$$

The Sharpe ratio. Given the expression for $\pi_{R,t}^U/\sigma_{R,t}^U$, we have

$$\begin{aligned} \frac{d(\pi_{R,t}^U/\sigma_{R,t}^U)}{d\theta_1} &= \frac{\sigma (\rho^h + \theta_1^h)^2 x_t}{[(\rho\theta_1^h - \rho^h\theta_1) x_t + \rho^h (1 + \theta_1/\rho)]^2} = \frac{\sigma\gamma^h x_t}{[\rho(\gamma^h - \gamma) x_t + \gamma]^2} > 0, \\ \frac{d(\pi_{R,t}^U/\sigma_{R,t}^U)}{d\theta_1^h} &= \frac{\sigma (1 + \theta_1/\rho) \rho^h (\theta_1 + \rho) \left(\frac{1}{\rho} - x_t\right)}{[(\rho\theta_1^h - \rho^h\theta_1) x_t + \rho^h (1 + \theta_1/\rho)]^2} \frac{\sigma\gamma^2 (1 - \rho x_t)}{\rho^h [\rho(\gamma^h - \gamma) x_t + \gamma]^2} \geq 0, \\ \frac{d(\pi_{R,t}^U/\sigma_{R,t}^U)}{d\bar{\theta}_1} &= \frac{\sigma (1 + \bar{\theta}_1/\rho) \left[\Delta\rho \left(\rho^h - \frac{\bar{\theta}_1^2}{\rho}\right) x_t + \rho^h (1 + \bar{\theta}_1/\rho)^2 \right]}{[-\Delta\rho\bar{\theta}_1 x_t + \rho^h (1 + \bar{\theta}_1/\rho)]^2}. \end{aligned}$$

When $\Delta\rho \left(\rho^h - \frac{\bar{\theta}_1^2}{\rho}\right) x_t + \rho^h (1 + \bar{\theta}_1/\rho)^2 > 0$, $\frac{d(\pi_{R,t}^U/\sigma_{R,t}^U)}{d\bar{\theta}_1} > 0$ and

$$\left\{ \begin{array}{l} \rho^h - \frac{\bar{\theta}_1^2}{\rho} = 0 \\ \left\{ \begin{array}{l} \rho^h - \frac{\bar{\theta}_1^2}{\rho} < 0 \\ x_t < -\frac{\rho^h(1+\bar{\theta}_1/\rho)^2}{\Delta\rho(\rho^h - \frac{\bar{\theta}_1^2}{\rho})} \end{array} \right. \end{array} \right. \Leftrightarrow \bar{\theta}_1 = \sqrt{\rho\rho^h}$$

Taking first derivative of ϕ with respect to $\bar{\theta}_1$ yields:

$$\frac{d\phi}{d\bar{\theta}_1} = \frac{-2\rho^h (1 + \bar{\theta}_1/\rho)}{\Delta\rho (\rho\rho^h - \bar{\theta}_1^2)^2} \left(\rho^h - \bar{\theta}_1 - 2\frac{\bar{\theta}_1^2}{\rho} \right).$$

Setting it to be zero gives:

$$\bar{\theta}_1 = \frac{-\rho \pm \sqrt{\rho^2 + 8\rho\rho^h}}{4}.$$

Since $\bar{\theta}_1 \geq 0$, we only consider the positive root. We now show $\bar{\theta}_1 = \frac{-\rho + \sqrt{\rho^2 + 8\rho\rho^h}}{4} < \sqrt{\rho\rho^h}$. Note that $\frac{\bar{\theta}_1}{\sqrt{\rho\rho^h}} = \frac{-\sqrt{\frac{\rho}{\rho^h}} + \sqrt{\frac{\rho}{\rho^h} + 8}}{4} < \frac{3}{4} < 1$, where we use the fact that $0 < \frac{\rho}{\rho^h} \leq 1$. Then we have $\rho^h - \bar{\theta}_1 - 2\frac{\bar{\theta}_1^2}{\rho} < 0$ for $\bar{\theta}_1\sqrt{\rho\rho^h}$, which implies that $\frac{d\phi}{d\theta_1} > 0$.

Now we show that $\phi > 1/\rho$. Given that $-\frac{\rho^h(1+\bar{\theta}_1/\rho)^2}{\Delta\rho\left(\rho^h - \frac{\bar{\theta}_1^2}{\rho}\right)} > \frac{1}{\rho}$, we have

$$-\rho\rho^h \left(1 + \frac{2\bar{\theta}_1}{\rho} + \frac{\bar{\theta}_1^2}{\rho}\right) > (\rho^h)^2 - \frac{\rho^h}{\rho}\bar{\theta}_1^2 - \rho\rho^h + \bar{\theta}_1^2 \text{ or } (\bar{\theta}_1 + \rho^h)^2 > 0,$$

which means that $x_t < \frac{1}{\rho} < \phi$. We then have $\frac{d(\pi_{R,t}^U/\sigma_{R,t}^U)}{d\theta_1} > 0$.

The Exposure price. Given that $\frac{dk_t}{d\theta_1} = \frac{\sigma^2(1+m)}{(m\rho^h + \rho)^2} \frac{\rho^h(1-\rho x_t - \rho m x_t)}{[1+(\rho^h - \rho)x_t]x_t}$ and

$$\bar{x}^c = \frac{1}{\tilde{m}\rho^h + \rho} = \frac{1}{\rho} \frac{1}{1 + \frac{\rho^h + \bar{\theta}_1}{\rho + \theta_1} m} \leq \frac{1}{\rho} \frac{1}{1 + m},$$

we have

$$x_t \leq \bar{x}^c \leq \frac{1}{\rho} \left(\frac{1}{1+m} \right) \text{ or } 1 - \rho x_t(1+m) > 0, \quad (1)$$

which means that

$$\frac{dk_t}{d\theta_1} = \frac{\sigma^2(1+m)}{(m\rho^h + \rho)^2} \frac{\rho^h [1 - \rho x_t(1+m)]}{[1 + (\rho^h - \rho)x_t]x_t} > 0.$$

The interest rate. Denote the Sharpe ratio $sp \equiv \frac{\pi_{R,t}}{\sigma_{R,t}}$, the expression for the interest rate can be written as:

$$r_t = \rho^h + \hat{g}_t - \rho(\rho^h - \rho)x_t - \rho q_t x_t - \frac{\rho x_t \left[(sp/\gamma)^2 - 2\sigma(sp/\gamma) \right] + \sigma^2}{1 - \rho x_t}.$$

Taking derivative with respect to θ_1 gives:

$$\frac{dr_t}{d\theta_1} = -\rho x_t \frac{dq_t}{d\theta_1} - \frac{2x_t}{(1 - \rho x_t)\gamma^2} \left[\left(\frac{sp}{\gamma} - \sigma \right) \left(-sp + \frac{dsp}{d\theta_1} \rho \gamma \right) \right], \quad (2)$$

which means that in the unconstrained case,

$$\frac{dr_t^U}{d\theta_1} = \frac{2\sigma^2 x_t \gamma^h (1 - \rho x_t)}{[\rho(\gamma^h - \gamma) x_t + \gamma]^3} (\gamma^h - \gamma) = \begin{cases} \frac{dr_t^U}{d\theta_1} < 0, & \text{if } \gamma^h < \gamma \\ \frac{dr_t^U}{d\theta_1} \geq 0, & \text{if } \gamma^h \geq \gamma. \end{cases}$$

Similarly, we have

$$\frac{dr_t}{d\theta_1^h} = -\rho x_t \frac{dq_t}{d\theta_1^h} - \frac{2\rho x_t}{(1 - \rho x_t)\gamma} \left(\frac{sp}{\gamma} \frac{dsp}{d\theta_1^h} - \sigma \frac{dsp}{d\theta_1^h} \right), \quad (3)$$

which means that

$$\frac{dr_t^U}{d\theta_1^h} = -\frac{2\sigma^2 \rho \gamma x_t (1 - \rho x_t) (\gamma^h - \gamma)}{\rho^h [\rho(\gamma^h - \gamma) x_t + \gamma]^3} \begin{cases} \frac{dr_t^U}{d\theta_1^h} > 0, & \text{if } \gamma^h < \gamma \\ \frac{dr_t^U}{d\theta_1^h} \leq 0, & \text{if } \gamma^h \geq \gamma. \end{cases}$$

In addition,

$$\begin{aligned} \frac{dr_t^U}{d\theta_1} &= \frac{2\sigma^2 \bar{\theta}_1 \Delta \rho x_t [\Delta \rho (\rho^h \bar{\theta}_1 + \rho \rho^h) x_t - \Delta \rho \rho^h (1 + \bar{\theta}_1/\rho)]}{\rho (1 + \bar{\theta}_1/\rho) [-\Delta \rho \bar{\theta}_1 x_t + \rho^h (1 + \bar{\theta}_1/\rho)]^3} \\ &= -\frac{2\sigma^2 \bar{\theta}_1 \rho^h (\Delta \rho)^2 x_t (1 - \rho x_t)}{\rho [-\Delta \rho \bar{\theta}_1 x_t + \rho^h (1 + \bar{\theta}_1/\rho)]^3} \leq 0. \end{aligned}$$

In the constrained case, from Equations (2) and (3), we have

$$\begin{aligned} \frac{dr_t}{d\theta_1} &= -\frac{\sigma^2 m}{(m\rho^h + \rho)^2 x_t} - \frac{2x_t}{(1 - \rho x_t)\gamma^2} \left[\left(\frac{\sigma}{(m\rho^h + \rho) x_t} - \sigma \right) \left(-\frac{\sigma \gamma}{(m\rho^h + \rho) x_t} + \frac{\sigma(\rho + \theta_1)}{\rho(m\rho^h + \rho) x_t} \right) \right] \\ &= -\frac{\sigma^2 m}{(m\rho^h + \rho)^2 x_t} < 0, \\ \frac{dr_t}{d\theta_1^h} &= \frac{\rho \sigma^2 m^2}{(m\rho^h + \rho)^2 (1 - \rho x_t)} > 0, \\ \frac{dr_t}{d\bar{\theta}_1} &= -\frac{\sigma^2 m [1 - \rho x_t (1 + m)]}{(m\rho^h + \rho)^2 (1 - \rho x_t) x_t} < 0 \end{aligned}$$

where we used the condition, (1).

Online Appendix B: Stationary Specialist Wealth Distribution

Without filtering, the stationary specialist wealth distribution can be solved explicitly. Specifically, let $G = 1 - \frac{\gamma^h}{\gamma} \in [0, 1]$ and $\Xi = \frac{\sigma \gamma^h}{1 - G\rho x}$, we have

$$\begin{aligned}\mu_u(x) &= \sigma^2 + \Delta\rho(1 - \rho x) - \frac{\rho x (\Xi^2 - 2\sigma\Xi) + \sigma^2}{1 - \rho x} + \Xi^2 - \sigma\Xi \\ &= \sigma^2 + \Delta\rho(1 - \rho x) + \sigma^2 \frac{(\gamma^h/\gamma)^2 (1 - 2\rho x) + (\gamma^h/\gamma) (1 - \rho Gx) (3\rho x - 1) - (1 - \rho Gx)^2}{(1 - \rho x) (1 - \rho Gx)^2}\end{aligned}$$

and

$$\sigma_u^2(x) = \sigma^2 G^2 \left(\frac{1 - \rho x}{1 - \rho Gx} \right)^2.$$

We can then compute that:

$$\begin{aligned}\int \frac{2\mu_u(x)}{\sigma_u^2(x)} dx &= \frac{2}{G^2} \left[\int \left(\frac{1 - \rho Gs}{1 - \rho s} \right)^2 ds + \frac{\Delta\rho}{\sigma^2} \int \frac{(1 - \rho Gs)^2}{1 - \rho s} ds + \left(\frac{\gamma^h}{\gamma} \right)^2 \int \frac{1 - 2\rho s}{(1 - \rho s)^3} ds \right. \\ &\quad \left. + \frac{\gamma^h}{\gamma} \int \frac{(1 - \rho Gs) (3\rho s - 1)}{(1 - \rho s)^3} ds - \int \frac{(1 - \rho Gs)^2}{(1 - \rho s)^3} ds \right] \\ &= \frac{2}{G^2} \left[\frac{\rho G^2 x (\rho x - 2) + 2(G - 1) G (\rho x - 1) \log(\rho x - 1) + 2G - 1}{\rho(\rho x - 1)} \right. \\ &\quad - \frac{\Delta\rho \rho Gx [\rho(Gx - 4) + 2G] + 2(\rho - G)^2 \log(1 - \rho x)}{2\rho^3} + \left(\frac{\gamma^h}{\gamma} \right)^2 \frac{3 - 4\rho x}{2\rho(\rho x - 1)^2} \\ &\quad \left. + \frac{\gamma^h G (4 - 5\rho x) + 3G(\rho x - 1)^2 \log(\rho x - 1) + 3\rho x - 2}{\rho(\rho x - 1)^2} + \frac{2G^2 \log(\rho x - 1) - \frac{(G-1)[G(4\rho x-3)-1]}{(\rho x-1)^2}}{2\rho} \right].\end{aligned}$$

Let $H = m\rho^h + \rho$, we have

$$\mu_c(x) = \sigma^2 + \Delta\rho(1 - \rho x) + \sigma^2 \frac{A_0 x^2 + A_1 x + A_2}{(1 - \rho x) (Hx)^2} \text{ and } \sigma_c^2(x) = \sigma^2 \left(\frac{1 - Hx}{Hx} \right)^2,$$

which implies that:

$$\begin{aligned}\frac{2\mu_c(x)}{\sigma_c^2(x)} &= 2 \left(\frac{Hx}{1 - Hx} \right)^2 + 2 \frac{\Delta\rho}{\sigma^2} (1 - \rho x) \left(\frac{Hx}{1 - Hx} \right)^2 \\ &\quad + \frac{2A_0 x^2}{(1 - \rho x) (1 - Hx)^2} + \frac{2A_1 x}{(1 - \rho x) (1 - Hx)^2} + \frac{2A_2}{(1 - \rho x) (1 - Hx)^2}\end{aligned}$$

and

$$\begin{aligned}
\int \frac{2\mu_c(x)}{\sigma_c^2(x)} dx &= \frac{2 \left[Hx + \frac{1}{1-Hx} + 2 \log(1-Hx) \right]}{H} \\
&- \frac{2\Delta\rho}{\sigma^2} \frac{H^2 \rho x^2 - 2Hx(H-2\rho) + \frac{2(H-\rho)}{Hx-1} + 2(3\rho-2H) \log(1-Hx)}{H^2} \\
&- 2A_0 \frac{\frac{H-\rho}{H^2(Hx-1)} + \frac{(\rho-2H) \log(1-Hx)}{H^2} + \frac{\log(1-\rho x)}{\rho}}{(H-\rho)^2} - 2A_1 \frac{\frac{H-\rho}{H(Hx-1)} - \log(1-Hx) + \log(1-\rho x)}{(H-\rho)^2} \\
&+ 2A_2 \frac{\rho(Hx-1) \log(1-Hx) + (\rho-H\rho x) \log(1-\rho x) - H + \rho}{(H-\rho)^2(Hx-1)}.
\end{aligned}$$