

# Online Appendix for “Long-run Consumption Risk and Asset Allocation under Recursive Utility and Rational Inattention”

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## 1. Deriving the Stochastic Properties of Consumption Dynamics

Taking unconditional variance on both sides of the expression for individual consumption dynamics,

$$\Delta c_{t+1}^* = \theta \left\{ \frac{\alpha^* u_{t+1}}{1 - ((1 - \theta) / \phi) \cdot L} + \left[ \zeta_{t+1} - \frac{(\theta / \phi) \zeta_t}{1 - ((1 - \theta) / \phi) \cdot L} \right] \right\}, \quad (1)$$

yields

$$\begin{aligned} \text{var} [\Delta c_t^*] &= \theta^2 \left\{ \frac{\alpha^{*2} \omega^2}{1 - (1 - \theta)^2 / \phi^2} + \left[ 1 + \frac{\theta^2 / \beta^2}{1 - (1 - \theta)^2 / \beta^2} \right] \omega_{\xi}^2 \right\} \\ &= \theta^2 \left\{ \frac{1}{1 - (1 - \theta)^2 / \phi^2} + \left[ \frac{1}{(1 - (1 - \theta) / \phi^2) \theta} - \frac{1}{1 - (1 - \theta)^2 / \phi^2} \right] \right\} \alpha^{*2} \omega^2 \\ &= \frac{\theta \phi^2}{\phi^2 + \theta - 1} \alpha^{*2} \omega^2. \end{aligned}$$

Note that in the absence of the endogenous noise shocks (i.e.,  $\zeta_t = 0$ ), we have

$$\text{var} [\Delta c_t^*] = \theta^2 \frac{\alpha^{*2} \omega^2}{1 - (1 - \theta)^2 / \phi^2} = \frac{\theta^2}{1 - (1 - \theta) / \phi^2} \alpha^{*2} \omega^2.$$

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Using (1), we can compute the first-order autocovariance of consumption growth as follows:

$$\begin{aligned}
\text{cov}(\Delta c_t^*, \Delta c_{t+1}^*) &= \text{cov} \left( \theta \left\{ \frac{\alpha^* u_t}{1 - ((1-\theta)/\phi) \cdot L} + \left[ \xi_t - \frac{(\theta/\phi) \xi_{t-1}}{1 - ((1-\theta)/\phi) \cdot L} \right] \right\}, \right. \\
&\quad \left. \theta \left\{ \frac{\alpha^* u_{t+1}}{1 - ((1-\theta)/\phi) \cdot L} + \left[ \xi_{t+1} - \frac{(\theta/\phi) \xi_t}{1 - ((1-\theta)/\phi) \cdot L} \right] \right\} \right) \\
&= \frac{1-\theta}{\phi} \text{cov} \left( \frac{\theta \alpha^* u_t}{1 - ((1-\theta)/\phi) \cdot L}, \frac{\theta \alpha^* u_t}{1 - ((1-\theta)/\phi) \cdot L} \right) \\
&\quad + \text{cov} \left( \theta \left[ \xi_t - \frac{(\theta/\phi) \xi_{t-1}}{1 - ((1-\theta)/\phi) \cdot L} \right], -\frac{\theta (\theta/\phi) \xi_t}{1 - ((1-\theta)/\phi) \cdot L} \right) \\
&= \frac{1-\theta}{\phi} \text{var} \left( \frac{\theta \alpha^* u_t}{1 - ((1-\theta)/\phi) \cdot L} \right) + \text{cov} \left( \theta \left[ \xi_t - \frac{(\theta/\phi) \xi_{t-1}}{1 - ((1-\theta)/\phi) \cdot L} \right], -\frac{\theta (\theta/\phi) \xi_t}{1 - ((1-\theta)/\phi) \cdot L} \right) \\
&= \frac{1-\theta}{\phi} \frac{(\theta \alpha^*)^2 \omega^2}{1 - (1-\theta)^2 / \phi^2} - \frac{\theta^3}{\phi} \frac{\alpha^{*2} \omega^2}{(1/(1-\theta) - 1/\phi^2) \theta} \\
&\quad + \frac{\theta^4 (1-\theta)}{\phi^3} \frac{1}{1 - (1-\theta)^2 / \phi^2} \frac{\alpha^{*2} \omega^2}{(1/(1-\theta) - 1/\phi^2) \theta} \\
&= \left[ \frac{1-\theta}{\phi} \frac{\theta^2}{1 - (1-\theta)^2 / \phi^2} - \frac{\theta^2}{\phi} \frac{1}{1/(1-\theta) - 1/\phi^2} + \frac{\theta^3 (1-\theta)}{\phi^3} \frac{1}{1 - (1-\theta)^2 / \phi^2} \frac{1}{1/(1-\theta) - 1/\phi^2} \right] \alpha^{*2} \omega^2 \\
&= 0.
\end{aligned}$$

We thus have

$$\text{corr}(\Delta c_t^*, \Delta c_{t+1}^*) = \frac{\text{cov}(\Delta c_t^*, \Delta c_{t+1}^*)}{\sqrt{\text{var}(\Delta c_t^*)} \sqrt{\text{var}(\Delta c_{t+1}^*)}} = 0.$$

Note that in the absence of the endogenous noise shocks, we have

$$\begin{aligned}
\text{corr}(\Delta c_t^*, \Delta c_{t+1}^*) &= \frac{1-\theta}{\phi} \frac{(\theta \alpha^*)^2 \omega^2}{1 - (1-\theta)^2 / \phi^2} \left( \frac{\theta \phi^2}{\phi^2 + \theta - 1} \alpha^{*2} \omega^2 \right)^{-1} \\
&= \frac{1-\theta}{\phi} \frac{\theta^2}{1 - (1-\theta)^2 / \phi^2} \left( \frac{\theta}{1 - (1-\theta)^2 / \phi^2} \right)^{-1} = \frac{\theta(1-\theta)}{\phi} > 0
\end{aligned}$$

because

$$\begin{aligned}
\text{cov}(\Delta c_t^*, \Delta c_{t+1}^*) &= \text{cov} \left( \frac{\theta \alpha^* u_t}{1 - ((1-\theta)/\phi) \cdot L}, \frac{\theta ((1-\theta)/\phi) \alpha^* u_t}{1 - ((1-\theta)/\phi) \cdot L} \right) \\
&= \frac{1-\theta}{\phi} \text{cov} \left( \frac{\theta \alpha^* u_t}{1 - ((1-\theta)/\phi) \cdot L}, \frac{\theta \alpha^* u_t}{1 - ((1-\theta)/\phi) \cdot L} \right) \\
&= \frac{1-\theta}{\phi} \frac{(\theta \alpha^*)^2 \omega^2}{1 - (1-\theta)^2 / \phi^2}.
\end{aligned}$$

Finally, using (1), it is straightforward to show that

$$\begin{aligned}\text{corr}(\Delta c_{t+1}^*, u_{t+1}) &= \frac{\text{cov}(\Delta c_{t+1}^*, u_{t+1})}{\text{sd}(\Delta c_{t+1}^*) \text{sd}(u_{t+1})} \\ &= \frac{\theta \alpha^* \omega^2}{\sqrt{\frac{\theta \phi^2}{\phi^2 + \theta - 1} \alpha^* \omega^2}} \\ &= \sqrt{\theta (1 - (1 - \theta) / \phi^2)},\end{aligned}$$

where we use the result that

$$\text{cov}(\Delta c_{t+1}^*, u_{t+1}) = \text{cov}(\theta \alpha^* u_{t+1}, u_{t+1}) = \theta \alpha^* \omega^2.$$

Note that in the absence of the endogenous noise shocks, we have

$$\text{corr}(\Delta c_{t+1}^*, u_{t+1}) = \sqrt{(1 - (1 - \theta) / \phi^2)},$$

## 2. Deriving the Long-run Risk in the Presence of the Correlated Noise

Substituting (1) into

$$\pi = \text{cov}_t \left[ \frac{\rho}{\sigma} \left( \sum_{j=0}^S \Delta c_{t+1+j} \right) + (1 - \rho) \left( \sum_{j=0}^S r_{p,t+1+j} \right), u_{t+1} \right], \quad (2)$$

we have

$$\begin{aligned}\pi &= \lim_{S \rightarrow \infty} \left\{ \sum_{j=0}^S \text{cov}_t \left[ \frac{\rho}{\sigma} \theta \left\{ \frac{\alpha u_{t+1+j}}{1 - ((1 - \theta) / \phi) \cdot L} + \left[ \zeta_{t+1+j} - \frac{(\theta / \phi) \zeta_{t+j}}{1 - ((1 - \theta) / \phi) \cdot L} \right] \right\} \right. \right. \\ &\quad \left. \left. + (1 - \rho) \left( \sum_{j=0}^S r_{p,t+1+j} \right), u_{t+1} \right] \right\} \\ &= \frac{\rho}{\sigma} \theta \left\{ \frac{1}{1 - (1 - \theta) / \phi} \alpha \omega^2 + \left[ 1 - \frac{\theta}{\phi} \frac{1}{1 - (1 - \theta) / \phi} \right] \rho \omega \omega_\zeta \right\} + (1 - \rho) \alpha \omega^2 \\ &= \frac{\rho}{\sigma} \theta \left\{ \frac{1}{1 - (1 - \theta) / \phi} + \rho u_\zeta \left[ 1 - \frac{\theta}{\phi} \frac{1}{1 - (1 - \theta) / \phi} \right] \sqrt{\frac{1}{[1 / (1 - \theta) - (1 / \phi)^2] \theta}} \right\} \alpha \omega^2 + (1 - \rho) \alpha \omega^2 \\ &= \left\{ \left[ \frac{\rho}{\sigma} \frac{\theta}{1 - (1 - \theta) / \phi} + (1 - \rho) \right] + \frac{\rho \rho u_\zeta \theta}{\sigma} \left[ 1 - \frac{\theta}{\phi} \frac{1}{1 - (1 - \theta) / \phi} \right] \sqrt{\frac{1}{[1 / (1 - \theta) - (1 / \phi)^2] \theta}} \right\} \alpha \omega^2,\end{aligned}$$

which will reduce to the expression for  $\Gamma$  in the text. Note that here we use the fact that  $\omega_\zeta = \sqrt{\frac{1}{[\exp(2\kappa) - (1 / \phi)^2] \theta}} \alpha \omega$ .

### 3. Deriving Optimal Consumption and Portfolio Rules in the Presence of Uninsurable Labor Income

Log-linearizing the flow budget constraint,  $A_{t+1} = R_{t+1}^p (A_t + Y_t - C_t)$ , around the long-run means of the log consumption-income ratio and the log wealth-income ratio,  $c - y = E [c_t - y_t]$  and  $a - y = E [a_t - y_t]$ , yields the approximate budget constraint

$$a_{t+1} - y_{t+1} = \rho_0 + \rho_a (a_t - y_t) + \rho_c (c_t - y_t) - \Delta y_{t+1} + r_{t+1}^p \quad (3)$$

where  $\rho$ ,  $\rho_a$ , and  $\rho_c$  are constants:

$$\begin{aligned} \rho_a &= \frac{\exp(a - y)}{1 + \exp(a - y) - \exp(c - y)} > 0, \\ \rho_c &= \frac{\exp(c - y)}{1 + \exp(a - y) - \exp(c - y)} > 0, \\ \rho_0 &= -(1 - \rho_a + \rho_c) \log(1 - \rho_a + \rho_c) - \rho_a \log(\rho_a) + \rho_c \log(\rho_c). \end{aligned}$$

Starting from the standard full-information rational expectations model with expected utility ( $\gamma = \sigma$  and  $\theta = 1$ ), we obtain the decision rules

$$\begin{aligned} c_t &= y_t + b_0 + b_1 (a_t - y_t) \\ \alpha^* &= \frac{1}{b_1} \left( \frac{\mu - r_f + \frac{1}{2}\omega^2}{\gamma\omega^2} \right) + \left( 1 - \frac{1}{b_1} \right) \frac{\omega_{uv}}{\omega^2} \end{aligned}$$

where

$$\begin{aligned} b_1 &= \frac{\rho_a - 1}{\rho_c} \in (0, 1], \\ b_0 &= \frac{1}{1 - \rho_a} \left[ \left( \frac{1}{\gamma} - b_1 \right) E [r_{t+1}^p] + \frac{1}{\gamma} \log(\beta) + \frac{\Xi}{2\gamma} - \rho_0 - (1 - b_1) g \right]. \end{aligned}$$

Here  $b_1$  is the elasticity of consumption with respect to financial wealth, making  $1 - b_1$  the elasticity with respect to labor income, and  $\Xi$  is an irrelevant constant term. If labor income is tradable,  $b_1 = 1$  and the model reduces to the one studied previously.

To help introduce rational inattention, we define a new state variable

$$s_t = a_t + \lambda y_t, \quad (4)$$

where

$$\lambda = \frac{1 - \rho_a + \rho_c}{\rho_a - 1}.$$

(As we have noted earlier, multivariate rational inattention models are analytically intractable, so the reduction of the state space to a single variable is critical for our results.) Using this new state

variable, the log-linearized budget constraint (3) can be rewritten as

$$s_{t+1} = \rho_0 + \rho_a s_t - \rho_c c_t - g + \rho_\varepsilon \varepsilon_{t+1} + \lambda v_{t+1} + r_{t+1}^p. \quad (5)$$

The consumption function can thus be rewritten as:

$$c_t = b_0 + b_1 s_t. \quad (6)$$

As in the benchmark model, applying the separation principle yields:

$$c_t = b_0 + b_1 \hat{s}_t \quad (7)$$

and we obtain the law of motion for the conditional mean of permanent income

$$\hat{s}_{t+1} = (1 - \theta) \hat{s}_t + \theta (s_{t+1} + \xi_{t+1}) + Y, \quad (8)$$

where  $Y$  is an irrelevant constant and all other notation is the same as that used in the benchmark model.

Given the assumption that  $1 - (1 - \theta) \rho_a > 0$ , applying the same long-term Euler equation we used in the benchmark model, we can solve for the optimal share invested in equity in the presence of labor income as:

$$\alpha^* = \frac{1}{\tilde{\zeta}} \left[ \frac{1}{b_1} \left( \frac{\mu - r^f + 0.5\omega^2}{\omega^2} \right) + \left( 1 - \frac{1}{b_1} \right) \frac{\tilde{\zeta} \omega_{uv}}{\omega^2} \right] \quad (9)$$

where  $\tilde{\zeta} = \frac{\rho}{\sigma} \zeta + 1 - \rho$  and  $\zeta = \frac{\theta}{1 - (1 - \theta) \rho_a} > 1$ .